CMPT 210: Probability and Computing

Lecture 25

Sharan Vaswani

April 11, 2024

Sample (outcome) space S: Nonempty (countable) set of possible outcomes.

Outcome $\omega \in \mathcal{S}$: Possible "thing" that can happen.

Event *E*: Any subset of the sample space.

Probability function on a sample space S is a total function $Pr : S \rightarrow [0, 1]$. For any $\omega \in S$,

$$0 \le \Pr[\omega] \le 1$$
 ; $\sum_{\omega \in S} \Pr[\omega] = 1$; $\Pr[E] = \sum_{\omega \in E} \Pr[\omega]$

Union: For mutually exclusive events E_1, E_2, \ldots, E_n , $\Pr[E_1 \cup E_2 \cup \ldots E_n] = \Pr[E_1] + \Pr[E_2] + \ldots + \Pr[E_n]$.

Complement rule: $\Pr[E] = 1 - \Pr[E^c]$

Inclusion-Exclusion rule: For any two events $E, F, \Pr[E \cup F] = \Pr[E] + \Pr[F] - \Pr[E \cap F]$. **Union Bound**: For any events $E_1, E_2, E_3, \dots, E_n, \Pr[E_1 \cup E_2 \cup E_3 \dots \cup E_n] \le \sum_{i=1}^n \Pr[E_i]$. **Uniform probability space**: A probability space is said to be uniform if $\Pr[\omega]$ is the same for

every outcome $\omega \in S$. In this case, $\Pr[E] = \frac{|E|}{|S|}$.

Conditional Probability: For events *E* and *F*, probability of event *E* conditioned on *F* is given by $\Pr[E|F]$ and can be computed as $\Pr[E|F] = \frac{\Pr[E \cap F]}{\Pr[F]}$.

Probability rules with conditioning: For the complement E^c , $\Pr[E^c|F] = 1 - \Pr[E|F]$.

Conditional Probability for multiple events: $\Pr[E_1 \cap E_2 \cap E_3] = \Pr[E_1] \Pr[E_2|E_1] \Pr[E_3|E_1 \cap E_2].$

Bayes rule: For events *E* and *F* if $Pr[E] \neq 0$, $Pr[F|E] = \frac{Pr[E|F]Pr[F]}{Pr[E]}$.

Law of Total Probability: For events *E* and *F*, $Pr[E] = Pr[E|F] Pr[F] + Pr[E|F^c] Pr[F^c]$.

Independent Events: Events *E* and *F* are said to be independent, if knowledge that *F* has occurred does not change the probability that *E* occurs, i.e. $\Pr[E|F] = \Pr[E]$ and $\Pr[E \cap F] = \Pr[E] \Pr[F]$.

Pairwise Independence: Events E_1, E_2, \ldots, E_n are pairwise independent, if for *every* pair of events E_i and E_j $(i \neq j)$, $\Pr[E_i|E_j] = \Pr[E_i]$ and $\Pr[E_i \cap E_j] = \Pr[E_i] \Pr[E_j]$.

Mutual Independence: Events E_1, E_2, \ldots, E_n are mutually independent, if for *every* subset of events, the probability that all the selected events occur equals the product of the probabilities of the selected events. Formally, for every subset $S \subseteq \{1, 2, \ldots, n\}$, $\Pr[\cap_{i \in S} E_i] = \prod_{i \in S} \Pr[E_i]$.

Random variable: A random "variable" R on a probability space is a total function whose domain is the sample space S. The codomain is denoted by V (usually a subset of the real numbers), meaning that $R: S \to V$.

Indicator Random Variables: An indicator random variable corresponding to an event E is denoted as \mathcal{I}_E and is defined such that for $\omega \in E$, $\mathcal{I}_E[\omega] = 1$ and for $\omega \notin E$, $\mathcal{I}_E[\omega] = 0$.

Probability density function (PDF): Let *R* be a random variable with codomain *V*. The probability density function of *R* is the function $PDF_R : V \to [0, 1]$, such that $PDF_R[x] = Pr[R = x]$ if $x \in Range(R)$ and equal to zero if $x \notin Range(R)$.

$$\sum_{x \in V} \mathsf{PDF}_R[x] = \sum_{x \in \mathsf{Range}(\mathsf{R})} \mathsf{Pr}[R = x] = 1.$$

Cumulative distribution function (CDF): The cumulative distribution function of *R* is the function $CDF_R : \mathbb{R} \to [0, 1]$, such that $CDF_R[x] = Pr[R \le x]$.

Distribution over a random variable can be fully specified using the cumulative distribution function (CDF) (usually denoted by F). The corresponding probability density function (PDF) is denoted by f.

Wrapping up

Bernoulli Distribution: $f_p(0) = 1 - p$, $f_p(1) = p$. Example: When tossing a coin such that Pr[heads] = p, random variable R is equal to 1 if we get a heads (and equal to 0 otherwise). In this case, R follows the Bernoulli distribution i.e. $R \sim Ber(p)$.

Uniform Distribution: If $R : S \to V$, then for all $v \in V$, f(v) = 1/|V|. *Example*: When throwing an *n*-sided die, random variable R is the number that comes up on the die. $V = \{1, 2, ..., n\}$. In this case, R follows the Uniform distribution i.e. $R \sim \text{Uniform}\{1, 2, ..., n\}$.

Binomial Distribution: $f_{n,p}(k) = {n \choose k} p^k (1-p)^{n-k}$. Example: When tossing *n* independent coins such that $\Pr[\text{heads}] = p$, random variable *R* is the number of heads in *n* coin tosses. In this case, *R* follows the Binomial distribution i.e. $R \sim \text{Bin}(n, p)$.

Geometric Distribution: $f_p(k) = (1-p)^{k-1}p$. Example: When repeatedly tossing a coin such that $\Pr[\text{heads}] = p$, random variable R is the number of tosses needed to get the first heads. In this case, R follows the Geometric distribution i.e. $R \sim \text{Geo}(p)$.

Expectation/mean of a random variable R is denoted by $\mathbb{E}[R]$ and "summarizes" its distribution. Formally, $\mathbb{E}[R] := \sum_{\omega \in S} \Pr[\omega] R[\omega]$

Alternate definition of expectation: $\mathbb{E}[R] = \sum_{x \in \text{Range}(R)} x \Pr[R = x].$

Expectation of transformed r.v's: For a random variable $X : S \to V$ and a function $g : V \to \mathbb{R}$, we define $\mathbb{E}[g(X)]$ as follows: $\mathbb{E}[g(X)] := \sum_{x \in \text{Range}(X)} g(x) \Pr[X = x]$

Linearity of Expectation: For *n* random variables R_1, R_2, \ldots, R_n and constants a_1, a_2, \ldots, a_n , b_1, b_2, \ldots, b_n , $\mathbb{E}\left[\sum_{i=1}^n a_i R_i + b_i\right] = \sum_{i=1}^n a_i \mathbb{E}[R_i] + b_i$.

Conditional Expectation: For random variable *R*, the expected value of *R* conditioned on an event A is given by $\mathbb{E}[R|A] = \sum_{x \in \text{Range}(R)} x \Pr[R = x|A]$

Law of Total Expectation: If *R* is a random variable $S \to V$ and events A_1, A_2, \ldots, A_n form a partition of the sample space, then, $\mathbb{E}[R] = \sum_i \mathbb{E}[R|A_i] \Pr[A_i]$.

Wrapping up

Independent random variables: We define two random variables R_1 and R_2 to be independent if for all $x_1 \in \text{Range}(R_1)$ and $x_2 \in \text{Range}(R_2)$, events $[R_1 = x_1]$ and $[R_2 = x_2]$ are independent. More formally,

$$\Pr[(R_1 = x_1) \cap (R_2 = x_2)] = \Pr[(R_1 = x_1)] \Pr[(R_2 = x_2)]$$

Independent random variables: Two random variables R_1 and R_2 are independent if for all $x_1 \in \text{Range}(R_1)$ and $x_2 \in \text{Range}(R_2)$,

$$\Pr[(R_1 = x_1) | (R_2 = x_2)] = \Pr[(R_1 = x_1)]$$

$$\Pr[(R_2 = x_2) | (R_1 = x_1)] = \Pr[(R_2 = x_2)]$$

Expectation of product of r.v's: For two r.v's R_1 and R_2 , $\mathbb{E}[R_1 R_2] = \sum_{x \in \text{Range}(R_1 R_2)} x \Pr[R_1 R_2 = x].$

Expectation of product of independent r.v's: For independent r.v's R_1 and R_2 , $\mathbb{E}[R_1 R_2] = \mathbb{E}[R_1] \mathbb{E}[R_2]$.

Joint distribution between r.v's X and Y can be specified by its joint PDF as follows: PDF_{X,Y}[x, y] = Pr[X = x \cap Y = y].

If X and Y are independent random variables, $PDF_{X,Y}[x, y] = PDF_X[x] PDF_Y[y]$.

Marginalization: We can obtain the distribution for each r.v. from the joint distribution by marginalizing over the other r.v's i.e. $PDF_X[x] = \sum_i PDF_{X,Y}[x, y_i]$.

Variance: Standard way to measure the deviation from the mean. For r.v. *X*, $Var[X] = \mathbb{E}[(X - \mathbb{E}[X])^2] = \sum_{x \in Range(X)} (x - \mu)^2 Pr[X = x]$ where $\mu := \mathbb{E}[X]$. **Alternate definition of variance**: $Var[X] = \mathbb{E}[X^2] - \mu^2 = \mathbb{E}[X^2] - (\mathbb{E}[X])^2$.

Standard Deviation: For r.v. X, the standard deviation of X is defined as $\sigma_X := \sqrt{\operatorname{Var}[X]} = \sqrt{\mathbb{E}[X^2] - (\mathbb{E}[X])^2}.$

Properties of variance: For constants *a*, *b* and r.v. *R*, $Var[aR + b] = a^2Var[R]$.

Pairwise Independence of r.v's: Random variables $R_1, R_2, R_3, ..., R_n$ are pairwise independent if for any pair R_i and R_j , for $x \in \text{Range}(R_i)$ and $y \in \text{Range}(R_j)$, $\Pr[(R_i = x) \cap (R_j = y)] = \Pr[R_i = x] \Pr[R_j = y].$

Linearity of variance for pairwise independent r.v's: If R_1, \ldots, R_n are pairwise independent, $Var[R_1 + R_2 + \ldots R_n] = \sum_{i=1}^n Var[R_i].$

Properties of variance: If R_1, \ldots, R_n are pairwise independent, for constants a_1, a_2, \ldots, a_n and b_1, b_2, \ldots, b_n , $Var[\sum_{i=1}^n a_i R_i + b_i] = \sum_{i=1}^n a_i^2 Var[R_i]$.

Covariance: For two random variables *R* and *S*, the covariance between *R* and *S* is defined as: $Cov[R, S] = \mathbb{E}[(R - \mathbb{E}[R])(S - \mathbb{E}[S])] = \mathbb{E}[RS] - \mathbb{E}[R] \mathbb{E}[S].$

Properties of covariance: If *R* and *S* are independent r.v's, $\mathbb{E}[RS] = \mathbb{E}[R]\mathbb{E}[S]$ and Cov[R, S] = 0. Cov[R, R] = Var[R]. Cov[R, S] = Cov[S, R].

Variance of sum of r.v's: For r.v's R_1, R_2, \ldots, R_n , Var $\left[\sum_{i=1}^n R_i\right] = \sum_{i=1}^n \operatorname{Var}[R_i] + 2 \sum_{1 \le i < j \le n} \operatorname{Cov}[R_i, R_j]$.

If R_i and R_j are pairwise independent, $Cov[R_i, R_j] = 0$ and $Var\left[\sum_{i=1}^n R_i\right] = \sum_{i=1}^n Var[R_i]$.

Correlation: For two r.v's R_1 and R_2 , the correlation between R_1 and R_2 is defined as $\operatorname{Corr}[R_1, R_2] = \frac{\operatorname{Cov}[R_1, R_2]}{\sqrt{\operatorname{Var}[R_1]\operatorname{Var}[R_2]}}$. $\operatorname{Corr}[R_1, R_2] \in [-1, 1]$ and indicates the strength of the relationship between R_1 and R_2 .

Bernoulli: If $R \sim \text{Bernoulli}(p)$, $\mathbb{E}[R] = p$ and Var[R] = p(1-p). **Uniform**: If $R \sim \text{Uniform}(\{v_1, \dots, v_n\})$, $\mathbb{E}[R] = \frac{v_1 + v_2 + \dots + v_n}{n}$ and $\text{Var}[R] = \frac{[v_1^2 + v_2^2 + \dots + v_n^2]}{n} - \left(\frac{[v_1 + v_2 + \dots + v_n]}{n}\right)^2$. **Binomial**: If $R \sim \text{Bin}(n, p)$, $\mathbb{E}[R] = np$ and Var[R] = np(1-p). **Geometric**: If $R \sim \text{Geo}(p)$, $\mathbb{E}[R] = \frac{1}{p}$ and $\text{Var}[R] = \frac{1-p}{p^2}$. Tail inequalities bound the probability that the r.v. takes a value much different from its mean.

Markov's Theorem: If X is a non-negative random variable, then for all x > 0, $\Pr[X \ge x] \le \frac{\mathbb{E}[X]}{x}$.

Chebyshev's Theorem: For a r.v. X and all x > 0, $\Pr[|X - \mathbb{E}[X]| \ge x] \le \frac{\operatorname{Var}[X]}{x^2}$.

Weak Law of Large Numbers: Let G_1, G_2, \ldots, G_n be pairwise independent variables with the same mean μ and (finite) standard deviation σ . Define $T_n := \frac{\sum_{i=1}^n G_i}{n}$, then for every $\epsilon > 0$, $\lim_{n\to\infty} \Pr[|T_n - \mu| \le \epsilon] = 1$.

Chernoff Bound: If T_1, T_2, \ldots, T_n are mutually independent r.v's such that $0 \le T_i \le 1$ for all *i*. If $T := \sum_{i=1}^n T_i$, for all $c \ge 1$ and $\beta(c) := c \ln(c) - c + 1$, $\Pr[T \ge c\mathbb{E}[T]] \le \exp(-\beta(c)\mathbb{E}[T])$. We have studied random variables that can take on discrete values – number of heads when tossing a coin, the number on a dice or the number of attempts to hit the bullsye in a dart game.

We have used these discrete distributions for designing randomized algorithms for verifying matrix multiplication, finding the maximum cut in graphs, randomized quickselect and voter poll.

In many applications, it is often more natural to model quantities as continuous random variables, for example, the amount of time it takes to transmit a message over a noisy channel or study the distribution of income in a population.

Continuous random variables are often used in distributed computing and for machine learning – fitting a model that can effectively explain the collected data.

STAT 271: Probability and Statistics for Computing Science

- Continuous random variables and distributions
- Sampling and Parameter estimation
- Linear Regression
- Hypothesis testing
- Analysis of Variance

Questions?